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Productivity and recovery of forage fish under climate change and fishing:
North Sea sandeel as a case study

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RUNNING HEAD: Productivity and recovery of forage fish
ABSTRACT

Forage fish occupy a central position in marine foodwebs worldwide by mediating the transfer of energy and organic matter from lower to higher trophic levels. The lesser sandeel (Ammodytes marinus) is one of the ecologically and economically most important forage fish species in the North-east Atlantic, acting as a key prey for larger fish and sea birds, as well as supporting a large commercial fishery. In this case study, we investigate the underlying factors affecting recruitment and how these in turn affect productivity of the North Sea sandeel using long-term data and modelling. Our results demonstrated how sandeel productivity in the central North Sea (Dogger Bank) depends on a combination of external and internal regulatory factors, including fishing and climate effects, as well as density dependence and food availability of the preferred zooplankton prey (Calanus finmarchicus and Temora longicornis). Furthermore, our model scenarios suggest that while fishing largely contributed to the abrupt stock decline during the late 1990s and the following period of low biomass, a complete recovery of the stock to the highly productive levels of the early 1980s would only be possible through changes in the surrounding ecosystem, involving lower temperatures and improved feeding conditions. To that end, we stress the need for ecosystem-based management accounting for multiple internal and external factors occurring within the broader context of the ecosystem in which forage fish species, such as sandeel, play an important and integral part.

KEY WORDS: population dynamics, recruitment, modelling, climate change, fisheries management
1. INTRODUCTION

Forage fish occupy a central position in marine food-webs worldwide (Cury et al., 2000; Chekley et al., 2009). By feeding on zooplankton and acting as a primary prey for large fish, marine birds and mammals, these fish species effectively mediate the transfer and flow of energy and organic matter from lower to higher trophic levels in marine ecosystems. Furthermore, these small, fast growing and typically short-lived species usually show highly variable population dynamics, both in terms of individual growth and recruitment (Checkley et al., 2009; MacKenzie et al., 2012; Lindegren and Checkley 2013; Bartolino et al., 2014). Since only a few species dominate these mid-trophic levels, important ecosystem functions and services (including both supporting and provisioning services such as food for naturally occurring predators and for human consumption) are highly sensitive to changes in the abundances of these species (Cury et al., 2000; Engelhard et al., 2014; Lindegren et al., 2016). Therefore, understanding the underlying factors regulating their large-scale fluctuations and population dynamics is utterly important.

Several studies have demonstrated clear top-down fishing effects on forage fish, usually exemplified by the dramatic collapses of commercially important small pelagic fish stocks of herring, sardines and anchovies (e.g., Dragesund et al., 1980; Coetzee et al., 2008; Freon et al., 2008). In addition, bottom-up forcing, mediated through direct physiological responses to environmental factors or indirect responses to climate-driven changes in the composition and availability of zooplankton prey has been shown to underlie large-scale fluctuations of forage fish, such as the alternating dominance of sardines and anchovies in upwelling systems (Baumgartner et al., 1992; Schwartzlose et al., 1999; Alheit and Niquen 2004; van der Lingen et al., 2006; Checkley et al., 2017). Although many studies have identified single drivers, the potential synergy of multiple
internal and external factors in regulating population dynamics of forage fish, as well as their relative importance throughout ontogeny is a largely unresolved issue (Hjermann et al., 2004; Lindegren et al., 2011; 2013). Given the ecological importance and high socio-economic value of forage fish worldwide, an ecosystem-based management approach, taking into consideration multiple external and internal regulating processes, is necessary to promote sustainable exploitation of these important marine resources and the diverse set of higher trophic level predators they support (McLeod and Leslie 2009).

The lesser sandeel (*Ammodytes marinus*) is one of the ecologically and economically most important forage fish species in the North-east Atlantic, acting as the preferred prey for predatory fish, marine mammals and birds, as well as supporting a large and lucrative fishery (Furness 2002; Greenstreet et al., 2006; Engelhard et al., 2014; Dickey-Collas et al., 2014). The industrial sandeel fishery began in the 1960s, when it mainly targeted sandeel populations in the North Sea. At its peak in the 1990s this largely unregulated fishery was removing a sandeel biomass amounting to ~1 million tonnes per year. In the early 2000s, the stock biomass declined rapidly and so did the landings and profitability of the fishery (Dickey-Collas et al., 2013; Engelhard et al., 2014; Fig. 1a, b). In order to promote stock recovery, the fishery has been under quota regulations for almost a decade, yet stock biomass has remained low and not returned to the productive levels of the 1980s (van Deurs et al., 2009; Lynam et al., 2013; Engelhard et al., 2014; ICES 2017). Lesser sandeel in the central North Sea spend most of the year buried within the bottom sediment (Wright et al., 2000; van Deurs et al., 2011). The growth period is confined to spring when they resurface and feed on copepods, such as the large *Calanus* sp. and the medium sized *Temora longicornis* (Macer 1966; van Deurs et al., 2013, 2014). Although a large number of factors have been proposed as drivers of the population dynamics of sandeel, including biotic processes (Arnott and Ruxton 2002a;
Frederiksen et al., 2007; Dickey-Collas et al., 2013; van Deurs et al., 2013; Eigaard et al., 2014), as well as climate forcing and fishery effects (Berntsen et al., 1994; Arnott and Ruxton 2002b; Frederiksen et al., 2007; van Deurs et al., 2009; Eliasen et al., 2011, van Deurs et al., 2014) the ultimate mechanisms preventing stock recovery are largely unknown, let alone quantified. In this study, we investigate the underlying factors regulating recruitment and productivity of sandeel in the central North Sea (Dogger bank) and assess the productivity and recovery potential of the stock under different climate and fishing scenarios using a coupled modelling approach. Understanding the combination and potential synergy of these internal and external factors is a prerequisite for development of ecosystem-based management practices necessary to promote sustainable exploitation of sandeel and other marine living resources worldwide (McLeod and Leslie 2009).

2. MATERIAL AND METHODS

2.1. Data availability

A number of variables characterising the local physical conditions, as well as regional ocean-atmospheric forcing, were collected (Table 1). The local climate conditions were represented by average sea surface temperatures (SST) at the Dogger Bank in each year and quarter, based on the Hadley centre observational data set available on a one-degree grid cell resolution (Fig. 1c), as well as mean annual sea bottom temperatures available from ICES (http://ecosystemdata.ices.dk/). In order to represent regional ocean-atmospheric forcing acting at larger spatial scales, we included annual averages of the North Atlantic Oscillation Index (NAO), representing the leading Empirical Orthogonal Function (EOF) of sea level pressure over the Atlantic sector (20°-80°N, 90°W-40°E). The number of recruits at age zero (“0-group”; R) and spawning stock biomass (SSB) estimates were derived from stock assessments for the years 1983-2015 of sandeel in the Dogger Bank (Fig.
1a, b), i.e., representing sandeel assessment area 1 (ICES 2017; see also Fig. S1). To account for
density-dependent effects on sandeel recruitment (Arnott & Ruxton 2002a), the number of one-year
old individuals was included. Finally, the abundances of the key zooplankton prey species Calanus
fincmarchicus and Temora longicornis (Fig. 1d) in the Dogger Bank area in each quarter were
available from the long-term monitoring of the Continuous Plankton Recorder (CPR; Batten et al.,
2003; Johns 2017).

2.2. Statistical analysis and recruitment modelling

Generalized Additive Models (GAMs; Hastie and Tibshirani, 1990; Wood 2006) were used to
examine the relationship between sandeel recruitment success (R/SSB) and the set of biotic and
abiotic variables chosen as possible predictors during model fitting and selection. The following
linearized Ricker formulations with log-transformed recruitment success estimates as responses
were used:

\[
\log(R/SSB) = a + s(\log(SSB)) + s(\log(N1)) + s(V) + \epsilon
\]

(1)

where \(a\) is the intercept, \(s\) the thin plate smoothing function, SSB the spawning stock biomass, N1
the number of one-year olds, \(V\) a number of selected climate predictors potentially affecting sandeel
recruitment success (Table 1 ) and \(\epsilon\) the error term. Although the number of regression splines is
optimized (and penalized) by the generalized cross validation criterion (GCV), the degrees of
freedom of the spline smoother function (\(s\)) was further constrained to three knots (k=3) to allow for
potential nonlinearities, but restrict flexibility during model fitting. Finally, we applied a model
reduction routine based on the GCV and partial \(F\)-tests to find the best possible set of predictors. To
limit the number of candidate models and avoid false positive results arising from multiple model
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testing we restricted the set of models to include covariates averaged for the same quarter or as
annual means (Table S2). In addition, we performed a cross validation analysis by fitting the final
model to a randomly selected subset of the data (Picard and Cook 1984), i.e., amounting to 75% of
the observations, and then assessing the predictive accuracy of the models by comparing the
observed values with the predicted recruitment estimates for the remaining subset. The cross-
validation analysis was repeated 1000 times (i.e., with a new set of random and serially independent
draws each time) in order to assess the range of uncertainty associated with the predictions.

2.3. Age-structured model

In order to simulate stock dynamics under different climate and fishing scenarios we applied a
standard age-structured cohort model based on available information and parameters derived from
recent stock assessments (ICES 2017). The simulated population dynamics are represented by
numbers-at-age (N) distributed among 5 age classes (from 0 to 4+), where the so-called plus group
includes all fish 4 years and older. The following formulation was used:

$$N_{a+1,t} = N_{a,t}e^{-(F_a,t+M_a,t)}$$

(2)

where $N_{a,t}$ are number-at-age $a$ in year $t$, $F_{a,t}$ and $M_{a,t}$ the fishing mortality and the natural mortality
at age $a$ in year $t$, respectively. The simulations were performed by estimating R/SSB based on the
final S-R model (Eq. 1; Table 3) and the observed values of each covariate in a given year, where
SSB was estimated as the sum of the adult population given by the proportion of mature fish in each
age and year and their corresponding numbers and mean weight-at-age (ICES 2017). In addition, a
stochastic element was included by adding Gaussian noise ($e$; resampled randomly from the
residuals of the S-R models) to account for unexplained sources of recruitment variability. After
having accounted for intrinsic processes (i.e., growth, maturation, and natural mortality), as well as external factors (i.e., fishing mortality) in the age-structured cohort models, the forward simulation loop is reiterated by estimating R/SSB in the following year. Similar to the available stock assessment model the simulation model applies half-year time steps, where recruitment occurs from the 1st to the 2nd part of the year.

2.4. Model simulations and scenario testing

To evaluate the relative importance of the various factors affecting sandeel recruitment and survival we used the model to hindcast the population dynamics over a period from 1996 to 2009 during which the SSB showed an abrupt decline and subsequently remained at low levels. The hindcast simulations comprise a control scenario where all input variables (i.e., SST, *C. finmarchicus*, *T. longicornis* and fishing mortalities at age) were kept at observed levels, as well as a set of “treatments” represented by the following scenarios: (i) reduced F to current levels from 2010 to 2015 (Fig. 1b), (ii) reduced SST to mean levels during the 1980s (Fig. 1c); (iii) introduced high abundances in *C. finmarchicus* occurring in 1998, 2001, 2004 and 2007 (corresponding to the observed peak in 1990; Fig. 1d); (iv) and all treatments above (i-iii) together. Furthermore, we performed multiple stochastic simulations (i.e., by randomly resampling and reintroducing residuals from the final GAM model) and estimated the probability of collapse as the percentage of simulations in which SSB falls below the agreed stock reference point below which recruitment is impaired (B_{lim}=160 000 tonnes) for each combination of SST (i.e., ranging from 0 to +2 °C relative to mean SST) and fishing, given by a scaling factor on mean fishing mortalities-at-age (F) ranging from 0 to 2 (i.e., representing a fishing closure and doubling of mean Fs, respectively). Probabilities were based on consecutive SSB values residing below B_{lim} for a period longer than one generation.
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(i.e., amounting to 4 years). The simulations were initialized at the estimated numbers at age in
2015 run for 30 years, and replicated 1000 times for each combination of ΔSST and F multipliers.
While observed SST values were used for model fitting and validation, surrogate time series of SST
were used as input during simulations. Since marine climate is generally positively autocorrelated
(Steele and Henderson 1984), we generated “red-shifted” noise accurately resembling the natural
variability of the observed SST time series by allowing the simulated SST time series to fluctuate
with the same mean, variance and degree of first-year autocorrelation as the observations
(Lindegren et al., 2010). The abundances of *C. finmarchicus* and *T. longicornis* were introduced in
the S-R model by stochastic resampling of observed values in each year due to the lack of a clear
auto-correlated signal. All statistical analyses were conducted using the R software, version 2.15.1
(www.r-project.org).

3. RESULTS

3.1. Drivers of recruitment success

After model fitting and selection the set of covariates retained within the final GAM included
spawning stock biomass (SSB), the number of one-year old sandeel (N1), spring sea surface
temperature (2nd quarter) and spring abundances of both *C. finmarchicus* and *T. longicornis* (Table
2; Table S2). Although fitting on variables averaged during winter (1st quarter) yielded a slightly
lower GCV (Table S2), the resulting model included a spurious and ecologically unrealistic
negative effect of prey availability (*T. longicornis*) on recruitment success (Fig. S2). The chosen
model, including covariates averaged over spring instead, demonstrated low GCV and a high degree
of explained deviance overall (75.2%). The final relationships between recruitment success,
spawning stock biomass (SSB) and the number of one-year old sandeel (N1) were represented by
non-linear or linearly decreasing functions, respectively (Fig. 2a, b), where in the former case the
negative effect on recruitment success is exacerbated at higher values of log(SSB). The functional
relationship between recruitment success and SST was best described by a negative linear
relationship (Fig. 2c), while the effect of C. finmarchicus and T. longicornis were linear and
positive (Fig. 2d, e). The final model explains the long-term dynamics and inter-annual variability
in recruitment success and hindcasted SSB (based on the age-structured model) well throughout the
period (Fig. 3a, b). Furthermore, the cross-validation routine demonstrated a high degree of
explained deviance for models fitted to a random and serially independent subset of the data, as
well as accuracy in predicting the remaining data (Fig. 4a). In addition, the explained deviance and
significance (p-values) of the model terms remained high when successively fitted and annually
updated with data from 1997 to 2015 (Fig. 4b). Model residuals were normally distributed and
temporally uncorrelated for recruitment and recruitment success (Fig. S3). As a complement to the
GAM approach we provide a number of alternative model formulations based on standard Ricker
models and generalized linear models including polynomial terms to account for non-linear
interactions (Table S2). While the GAM formulation demonstrates the highest degree of explained
deviance, the alternative model formulations demonstrate that the explanatory variables identified in
the GAM approach contribute significantly to the prediction of recruitment and recruitment success
independent of model choice. This supports the robustness of our findings and additionally provides
parameterized models for future management applications.

3.2. Testing the relative importance of drivers through scenario simulations

The hindcast model simulations used to evaluate the relative importance of the various factors
affecting sandeel recruitment and survival showed that the scenario based on reduced fishing
mortalities lead to a pronounced improvement in stock status relative to the control simulation with SSB values rising above $B_{lim}$ throughout most of the period (Fig. 5a). The scenario of reduced SST demonstrates a more moderate improvement in SSB compared to the control but show SSB values above $B_{lim}$ in the early and late 2000s. The scenario introducing recurring high abundances in C. finmarchicus shows a rather minor response. Interestingly, the combined scenario introducing all “treatments” shows considerably higher SSB values well above $B_{lim}$. Finally, the model simulations of sandeel dynamics under different combinations of SST and exploitation illustrate a strong dependence on both factors with a high probability of collapse at high levels of F and SST (Fig. 5b).

4. DISCUSSION

In this study we have demonstrated how sandeel productivity and the potential for stock recovery in the Dogger Bank area of the North Sea depend on a combination of internal density-dependent processes and external environmentally-driven factors. Among the external factors considered in this study, SST in spring explained a large part of recruitment variability and demonstrated a clear negative relationship. Consequently, the pronounced long-term increase and unprecedented high SST values since the early 2000s can partly explain the poor recruitment success and lack of recovery during recent decades. Furthermore, we show a strong negative relationship between the abundance of one-year old sandeel and recruitment success (settling of young of the year fish) indicating pronounced density-dependent regulation, likely acting through food or habitat competition. The dependence on food availability is supported by the positive effect of C. finmarchicus and T. longicornis on recruitment success, illustrated by the close correspondence between peaks in prey abundance with years of high recruitment success. The set of drivers identified in our study largely corresponds with previous work on sandeel recruitment (Arnott and
Ruxton 2002; van Deurs et al., 2009). This indicates that environmental correlates in relation to this particular stock appear to be exceptionally robust over time, even when performed on significantly extended time series and despite a common tendency for published recruitment-environment correlations to fail when verified upon retest (Myers 1998). Based on this consistency and the strong support for the drivers and relationships identified we used a scenario-based modelling approach for assessing how different climate regimes and fishing pressures interact to determine stock productivity and the potential for stock recovery (Lehodey et al., 2006).

However, before discussing these scenarios, we address the causality between the set of drivers and their associated effects on recruitment and stock productivity.

First of all, increasing water temperature affects fish metabolism, resulting in higher growth rates, if food is not a limiting factor and/or if the temperature is not already near the thermal tolerance limit of the species (Pörtner 2002). In contrast, if food is limiting, growth will decline and the fish may starve because of increased energetic costs of metabolism. This effect will depend on temperature as a specific amount of food may be sufficient at low temperatures but cause decreased growth at high temperatures. Lesser sandeel inhabiting the Dogger area in the North Sea, are not living on the southern or the northern limit of the distribution range of this species (Jensen et al., 2010; Fishbase 2016). Consequently, the temperatures they experience do not appear to be at, or close to, its thermal limits. However, the combination of increasing metabolic demands, driven by increasing temperatures, together with stable or decreasing prey availability (i.e., abundance of C. finmarchicus and T. longicornis) suggest that food limitation can limit recruitment success and stock recovery. This is particularly important during late winter and spring as this period covers the larval phase (Wright and Bailey 1996). Increasing temperatures lead to increasing energy demands potentially affecting developmental rate of the larvae (Pankhurst and Munday 2011). Furthermore,
in late winter and early spring food is exceptionally scarce and eggs and nauplii of \textit{C. finmarchicus}, advected into the North Sea from the Atlantic continental shelf, may represent a vital resource for the newly hatched larvae in this area of the North Sea (Arnott and Ruxton 2002; Richardson \textit{et al.}, 2011). In addition, \textit{T. longicornis} and other medium sized calanoid copepods often serve as other important prey in this area (Macer 1966; van Deurs \textit{et al.} 2014). These copepod species become abundant in late spring, around the time of sandeel metamorphosis, which is another critical life stage of sandeel (Wright and Bailey 1996). Hence, delayed or suppressed copepod production during this period may cause starvation and consequently limit survival and year class strength.

Although \textit{T. longicornis} is far more abundant than \textit{C. finmarchicus} in late spring (Pitois and Fox 2006), this less abundant prey species may still be important, owing to its larger size and higher fat content (van Deurs \textit{et al.} 2014).

While these mechanisms stress a strong link between stock productivity and the environment, the long-term increase in SST and decline of \textit{C. finmarchicus} and \textit{T. longicornis} in the North Sea should be understood within the broader context of large-scale ecosystem changes acting across the North East Atlantic (Beaugrand \textit{et al.}, 2004; Möllmann \textit{et al.}, 2009; Lindegren \textit{et al.}, 2012). These abrupt changes, often referred to as regime shifts, involve a number of additional factors potentially affecting the population dynamics of sandeel in the North Sea, such as an increasing number of competing forage fish species and predators on sandeel larvae and juveniles (Raab \textit{et al.}, 2012; Petitgas \textit{et al.}, 2012; Engelhard \textit{et al.}, 2014) and pronounced shifts in the phyto- and zooplankton community composition (Reid \textit{et al.}, 1998; Beaugrand 2004; Beare \textit{et al.}, 2004; Jansen and Gislason 2011). These effects are exacerbated by the resident nature of sandeel, which due to its strong association to a well-defined and patchily distributed habitat (Wright \textit{et al.}, 2000; Jensen \textit{et al.}).
al., 2011) cannot change distribution and take advantage of better conditions elsewhere (Helaouët and Beaugrand 2007).

After demonstrating the importance of both internal, density-dependent mechanisms and external, climate-driven factors, including temperature and food availability in regulating sandeel recruitment (Berntsen et al., 1994; Arnott and Ruxton 2002a; Frederiksen et al., 2007; Dickey-Collas et al., 2013; van Deurs et al., 2009; 2013; 2014; Eigaard et al., 2014), the question remains to what extent past changes in these drivers explain the poor productivity and virtual lack of recovery in recent decades. Our scenario-based modelling approach serves to provide insight into this issue by assessing the relative importance of these factors, particularly in relation to commercial fishing. Our results show that fishing largely contributed to the abrupt stock decline during the late 1990s and the following period of low biomass, especially given the extreme fishing mortalities experienced during this period (i.e. mean F=1.2 between 1999 to 2004; Fig. 1b). However, although reducing F to the currently lower levels would likely have maintained the stock biomass above ecologically safe levels (Fig. 5a) a complete recovery of the stock to the highly productive levels of the early 1980s would not have been possible, despite the lower level of exploitation. This highlights that factors besides fishing also contributed to the stock decline and to maintaining the stock in a state of poor productivity. We argue that these factors can be explained by the pronounced increase in SST, the concurrent decline in food availability, as well as the recent increase in natural mortality, especially through predation from mackerel. Whether recruitment and recovery of other sandeel stocks in the area are influenced or limited by the same set of drivers and processes is unclear and merit further attention. However, the joint dependence of recruitment and survival on fishing and environmental factors, including both abiotic and biotic processes, has been shown for several commercially important forage fish species (Freon et al., 2008; Lindegren et al., 2011, 2013;
Checkley et al., 2017), as well as for large predatory fish stocks worldwide, notably cod (Lindegren et al., 2009; Gårdmark et al., 2013; Pershing et al., 2015). This supports the generality of our findings and illustrates the need to adopt a more holistic ecosystem-based approach to fisheries management, accounting for multiple ecological and physical factors to promote sustainable exploitation of our living marine resources (MacLeod and Leslie 2009; Long et al., 2015). Despite a growing body of research on the effect of climate and biotic interactions on fish population dynamics, only a very limited number of fish stocks worldwide are currently accounting for these issues in tactical management (Skern-Mauritzen et al., 2015). To that end, we stress the need for ecosystem-based management accounting for multiple internal and external factors occurring within the broader context of the ecosystem, in which forage fish species, such as sandeel, play an important and integral part.

ACKNOWLEDGEMENTS

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REFERENCES


TABLES

**Table 1.** Information about the abiotic and biotic covariates used during model fitting with regards to spatial and temporal coverage, as well as the source of data.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Month</th>
<th>Area</th>
<th>Data source</th>
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</thead>
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<td>SSTQ1</td>
<td>Jan-March</td>
<td>Dogger Bank</td>
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</tr>
<tr>
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<td>April-June</td>
<td>Dogger Bank</td>
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</tr>
<tr>
<td>SST Annual</td>
<td>Jan-Dec</td>
<td>Dogger Bank</td>
<td><a href="http://www.metoffice.gov.uk/hadobs/hadisst/data/download.html">http://www.metoffice.gov.uk/hadobs/hadisst/data/download.html</a></td>
</tr>
<tr>
<td>SBT Annual</td>
<td>Jan-Dec</td>
<td>North Sea</td>
<td><a href="http://ecosystemdata.ices.dk/">http://ecosystemdata.ices.dk/</a></td>
</tr>
<tr>
<td><em>C. finmarchicus</em></td>
<td>Jan-Dec</td>
<td>Dogger Bank</td>
<td>Johns (2017)</td>
</tr>
<tr>
<td><em>T. longicornis</em></td>
<td>Jan-Dec</td>
<td>Dogger Bank</td>
<td>Johns (2017)</td>
</tr>
</tbody>
</table>
Table 2. Summary statistics of parametric coefficients and smooth terms for the final stock-recruitment model for North Sea sandeel.

A. Intercept

<table>
<thead>
<tr>
<th>Estimate</th>
<th>SE</th>
<th>t-value</th>
<th>p-value</th>
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</thead>
<tbody>
<tr>
<td>-0.52</td>
<td>0.1</td>
<td>-5.34</td>
<td>&lt;0.001***</td>
</tr>
</tbody>
</table>

B. Smooth terms

<table>
<thead>
<tr>
<th>Predictor</th>
<th>edf</th>
<th>F-value</th>
<th>p-value</th>
<th>Partial $r^2$ (%)</th>
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</thead>
<tbody>
<tr>
<td>SSB</td>
<td>1.71</td>
<td>33.1</td>
<td>&lt;0.001***</td>
<td>52.7</td>
</tr>
<tr>
<td>N1</td>
<td>1.00</td>
<td>19.8</td>
<td>&lt;0.001***</td>
<td>18.2</td>
</tr>
<tr>
<td>SST</td>
<td>1.00</td>
<td>10.0</td>
<td>0.004**</td>
<td>11.3</td>
</tr>
<tr>
<td>C. finmarchicus</td>
<td>1.25</td>
<td>4.53</td>
<td>0.031*</td>
<td>6.5</td>
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<tr>
<td>T. longocornis</td>
<td>1.00</td>
<td>5.89</td>
<td>0.023*</td>
<td>5.7</td>
</tr>
</tbody>
</table>

* edf is the estimated degrees of freedom for the model smooth terms where edf>1 indicates a non-linear relationship. The partial $r^2$ refer to the percentage of the total deviance explained by each covariate separately.
FIGURES CAPTIONS

Figure 1. Long-term trends in (a) sandeel recruitment (grey bars), spawning stock biomass (SSB; black line) and (b) mean fishing mortalities (F at ages 1-2; grey) estimated from stock assessments (ICES 2017). Horizontal dotted lines represent the precautionary and limiting stock sizes (B_{pa} and B_{lim}). Abiotic and biotic conditions affecting are recruitment success and juvenile survival are represented by (c) mean sea surface temperatures (SST) and the abundance of C. finmarchicus and T. longicornis at the Dogger Bank in spring (April-June) (d).

Figure 2. The partial effects of final model predictors on sandeel recruitment success with 95% confidence intervals (grey), illustrating negative relationships between log-recruitment success and spawning stock biomass (SSB) (a), the number of one-year old sandeel (b) and spring sea surface temperature (c), as well as positive effects of prey abundance during spring by C. finmarchicus (d) and T. longicornis (e).

Figure 3. (a) Observed (circles) and fitted values (black) of recruitment success with 95% confidence intervals (grey) based on the final GAM. (b) Observed (circles) and hindcasted estimates of spawning stock biomass (SSB; black) with 95% confidence intervals (grey) based on an age-structured population model.

Figure 4. (a) Boxplots of explained deviance from a cross-validation analysis of model fit on a randomly selected subset, as well as the associated accuracy of predictions on the remaining data (after 1000 model iterations). (b) Overall explained deviance (grey) and p-values of the effects of
SSB (solid), abundance at age 1 (dashed) and temperature (dotted) on sandeel recruitment success when successively fitting and annually updating the final model on data from 1997 to 2015. (The horizontal gray dashed line indicate significance at p=0.05).

**Figure 5.** (a) Hindcast simulations of sandeel SSB based on: (i) observed SST, *C. finmarchicus*, *T. longicornis* and F values (black); (ii) reduced F to current levels (2010-2015) (black dashed); (iii) reduced SST to mean levels during the 1980s (black dotted); (iv) introduced peaks in *C. finmarchicus* in 1998, 2001, 2004 and 2007 (grey solid); (v) and all changes (ii-iv) together (grey dashed). The solid horizontal red line marks the limiting stock level (B_{lim}). (b) Probability of SSB falling below B_{lim} given changes in mean SST (by -1 to 2°C) and fishing, given by a scaling factor on mean fishing mortalities-at-age (F multiplier) ranging from 0 to 2, hence representing a fishing closure and doubling of mean F, respectively (where 1 represent F equal to the long-term average, F_{age1-2} = 0.71).
Figure 1

218x209mm (72 x 72 DPI)
Figure 2

185x212mm (72 x 72 DPI)
Figure 3

161x203mm (72 x 72 DPI)
Figure 4

187x217mm (72 x 72 DPI)
Figure 5

254x190mm (96 x 96 DPI)
### SUPPLEMENTARY TABLES

**Table S1.** The generalized cross validation scores (GCV) and Akaike’s Information Criterion (AIC) after fitting the S-R model to each abiotic covariate and the zooplankton variables (*Calanus finmarchicus* and *Temora longicornis*) in the corresponding quarter (or as annual averages). The selected model is highlighted in bold.

<table>
<thead>
<tr>
<th>Variable</th>
<th>GCV</th>
<th>AIC</th>
</tr>
</thead>
<tbody>
<tr>
<td>SSTQ1*</td>
<td>0.380</td>
<td>62.6</td>
</tr>
<tr>
<td>SSTQ2</td>
<td>0.411</td>
<td>66.7</td>
</tr>
<tr>
<td>SSTQ3</td>
<td>0.532</td>
<td>74.8</td>
</tr>
<tr>
<td>SSTQ4</td>
<td>0.732</td>
<td>85.8</td>
</tr>
<tr>
<td>SSTAnnual</td>
<td>0.443</td>
<td>69.5</td>
</tr>
<tr>
<td>SBTAnnual</td>
<td>0.456</td>
<td>70.1</td>
</tr>
<tr>
<td>NAOAnnual</td>
<td>0.475</td>
<td>71.5</td>
</tr>
</tbody>
</table>

*Note that fitting the model on SST and zooplankton in quarter 1 yielded the lowest AIC/GCV. However, the model showed a spurious and ecologically unrealistic negative effect of prey availability (*Temora longicornis*) on recruitment success (Fig. S2) and was therefore excluded.*
Table S2. Summary of three alternative model parametric model formulations. Extended Ricker
\[
\ln(R/S) = c_1 S + c_2 \ln(N) + c_3 T + c_4 Z_1 + c_5 Z_2 + k \quad (M\ 1),
\]
simple linear regression 
\[
R = c_1 S + c_2 N + c_4 T + c_5 Z_1 + c_6 Z_2 + k \quad (M\ 2),
\]
and second order polynomial 
\[
\ln(R/S) = c_1 \ln(S) + c_2 \ln(S)^2 + c_3 \ln(N) + c_4 \ln(N)^2 + c_5 T + c_6 Z_1 + c_7 Z_2 + k \quad (M\ 3).
\]

The models use the same input as the GAM model in the main text of this manuscript: R (recruitment, in thousands individuals), S (spawning stock biomass, in tons), N (number of age-1 sandeel, in thousands individuals), T (sea surface temperature in °C), Z\(_1\) (CPR index for *Calanus finmarchicus*), and Z\(_2\) (CPR index for *Temora longicornis*). See material and method and Table 1 in the main text for a more detailed description of data sources. The table contains estimated coefficients (c\(_1\) – c\(_7\) and k) and the associated p-values (presented within parentheses, alpha = 0.05).

### Estimated coefficients:

<table>
<thead>
<tr>
<th>Model</th>
<th>c1</th>
<th>c2</th>
<th>c3</th>
<th>c4</th>
<th>c5</th>
<th>c6</th>
<th>c7</th>
<th>K</th>
<th>adj.R(^2)</th>
</tr>
</thead>
<tbody>
<tr>
<td>M1</td>
<td>-0.0000046</td>
<td>-0.72</td>
<td>-0.42</td>
<td>-0.22</td>
<td>-0.83</td>
<td></td>
<td></td>
<td>15.43</td>
<td>0.69</td>
</tr>
<tr>
<td>M2</td>
<td>-230.5</td>
<td>-843</td>
<td>-74150000</td>
<td>6.356</td>
<td>110800000</td>
<td></td>
<td></td>
<td>421500000</td>
<td>0.36</td>
</tr>
<tr>
<td>M3</td>
<td>9.46</td>
<td>-0.44</td>
<td>1.16</td>
<td>-0.085</td>
<td>-0.47</td>
<td>0.22</td>
<td>0.74</td>
<td>-44.45</td>
<td>0.71</td>
</tr>
</tbody>
</table>

### P-values for each coefficient:

<table>
<thead>
<tr>
<th>Model</th>
<th>c1</th>
<th>c2</th>
<th>c3</th>
<th>c4</th>
<th>c5</th>
<th>c6</th>
<th>c7</th>
<th>K</th>
</tr>
</thead>
<tbody>
<tr>
<td>M1</td>
<td>&lt;0.001***</td>
<td>&lt;0.001***</td>
<td>&lt;0.01**</td>
<td>0.042*</td>
<td>0.01*</td>
<td></td>
<td></td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>M2</td>
<td>0.112</td>
<td>0.025*</td>
<td>0.041*</td>
<td>0.017*</td>
<td>0.1</td>
<td></td>
<td></td>
<td>0.334</td>
</tr>
<tr>
<td>M3</td>
<td>0.084</td>
<td>0.047*</td>
<td>0.725</td>
<td>0.569</td>
<td>&lt;0.01**</td>
<td>0.072</td>
<td>0.018*</td>
<td>0.298</td>
</tr>
</tbody>
</table>
Figure S1. Map of the study area. Black round symbols illustrate the relative distributions of sandeel landings since 2003. Box with dashed boundaries envelopes Dogger bank and all landings from sandeel management area 1 (the Dogger bank population).
Figure S2. The partial effect of *Temora longicornis* in quarter 1 on sandeel recruitment success.

The model shows a spurious and ecologically unrealistic negative effect of prey availability (upper panel) which is entirely driven by three observations with low zooplankton abundances. If excluding these during fitting the relationship becomes insignificant (p=0.49) (lower panel).
Figure S3. (A) Normal probability plots and (B) partial autocorrelation plots of the final S-R models for sandeel recruitment success.