



Introducing selfisher: open source software for statistical analyses of fishing gear selectivity

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1 Introducing `selfisher`: open source
2 software for statistical analyses of
3 fishing gear selectivity

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17 Running Headline: R package `selfisher` for gear selectivity

18

19 Abstract

20 There is a need to improve fishing methods to select for certain sizes and species while
21 excluding others. Experiments are conducted to quantify selectivity of fishing gears and how
22 variables such as gear design (e.g. mesh size, mesh shape), environmental parameters (e.g.
23 light, turbidity, substrate) or biological parameters (e.g. fish condition) alter selectivity; the
24 resulting data need to be analyzed using specialized statistical methods in many cases. Here,
25 we present a new tool for analyzing this type of data: an R package named `selfisher`. It
26 allows estimating multiple fixed effects (e.g. fish length, total catch weight, environmental
27 variables) and random effects (e.g. haul). A bootstrapping procedure is also provided. We
28 demonstrate its use via four case studies including (A) covered codend analyses of four gears,
29 (B) a paired gear study with numerous covariates, (C) a catch comparison study of unpaired
30 hauls of gillnets and (D) a catch comparison study of paired hauls using polynomials and
31 splines. This software will make it easier to model selectivity, teach statistical methods, and
32 make analyses more repeatable.

33
34 **Keywords:** geometric similarity, catch comparison, covered codend, gillnet, mesh size, paired
35 gear, trawl

36
37

38 Introduction

39 The wasteful practice of returning organisms back to the sea, often dead or dying, is
40 commonly known as discarding and is responsible for approximately 11% of total commercial
41 catches or 9.1 million tons (Pérez Roda et al., 2019). Consequently, substantial effort is spent
42 on trying to reduce these unwanted catches by developing fishing gears that select for certain
43 species and sizes of individuals while allowing others to avoid capture (Kennelly and
44 Broadhurst, 2021). Experiments are conducted to measure the selectivity of fishing gear and
45 statistical models are used to characterize the selectivity patterns. The selectivity of fishing gear
46 is commonly described by a retention curve, i.e. the probability of being retained in the net,
47 which is usually a function of individual length or size and may vary between hauls (Wileman et
48 al., 1996). Between-haul variation may be random (due to unmeasured variables), or it may
49 depend on observed covariates such as total catch weight (Fryer, 1991; Suuronen & Millar,
50 1992; Erickson et al., 1996; O'Neill & Kynoch, 1996), environmental variables that affect animal
51 behavior including light, temperature, and bottom substrate (He, 1993; Walsh & Hickey, 1993;
52 Ryer & Barnett, 2006, Somerton et al., 2013), or the body condition of individuals (Özbilgin et
53 al., 2007; Ferro et al., 2008).

54 In some cases, catch data collected in selectivity studies could be analyzed with
55 standard binomial generalized linear models, for which there are plenty of software options
56 available. Binomial models are appropriate because, in many gear selectivity experiments,
57 individuals end up in one of two compartments (e.g. codend vs cover; gear 1 vs gear 2; or test
58 gear vs control gear; Wileman et al., 1996), i.e. there are two possible outcomes as in coin flips.
59 However, a substantial amount of the analyses in this field are specialized and require
60 specialized software. For example, obtaining confidence intervals on predictions typically
61 requires accounting for extra-binomial variability (overdispersion) between and within hauls

62 (Millar, 1993; Millar et al., 2004). Paired gear studies (where the test gear is tested against a
63 control one with retention probability equal to one for the given species and lengths of interest)
64 is another example that does not conform to a typical logistic regression model because the
65 probability model is more complex as we will show below. This paper presents newly developed
66 open source software that is specifically designed for modelling fishing gear selectivity,
67 something that was previously limited. The package, `selfisher`, is implemented in the R
68 statistical computing environment which is commonly used for many modern fisheries analyses
69 (R core team 2020). The package is written with an interface that will be familiar to many users
70 of regression methods in R. By making the software free and openly available, we aim to
71 improve repeatability of analyses and enable teaching these analytical methods in classrooms
72 for the next generation of fisheries scientists.

73 In this paper, first, we describe the implementation of the `selfisher` R package and
74 provide a general description of how models are estimated by the package, including dealing
75 with subsampled catches, a common occurrence in gear selectivity studies (e.g. Larsen et al.,
76 2018; Melli et al., 2019; Veiga-Malta et al., 2020). We also describe ways that the package can
77 be used to address the common issue of overdispersion (Millar et al., 2004). We describe three
78 general categories of statistical models, divided based on the mathematical probabilities
79 underlying the estimation in our method, while omitting details about experimental designs and
80 code to run the models. Then, we describe the bootstrapping procedure used to account for
81 variation within and between hauls in selectivity (Millar, 1993), potentially resampling from
82 distinct pools of hauls based on gear type or tactic as in unpaired hauls (Herrmann et al., 2017;
83 Savina et al., 2017). Then, we briefly describe four case studies used to illustrate the package's
84 capabilities, while providing more thorough descriptions with R code in the supplementary
85 material. The case studies are (A) a covered codend study of haddock (*Melanogrammus*
86 *aeglefinus*) with four different codends (O'Neill et al., 2016), (B) a paired gear study in a brown

87 shrimp (*Crangon crangon*) trawl fishery where one trawl is nonselective (Santos et al., 2018),
88 (C) a catch comparison study of unpaired hauls of gillnets avoiding an unwanted crab (*Cancer*
89 *pagurus*) (Savina et al., 2017), and (D) a catch comparison study of paired hauls in a Norway
90 lobster (*Nephrops norvegicus*) trawl fishery (Melli et al., 2018). Finally, we discuss current
91 limitations of the `selfisher` package and potential future advancements.

92 Implementation of `selfisher`

93 We designed the `selfisher` package to be flexible and robust for fitting and assessing
94 a variety of gear selectivity models that can be represented with a binomial distribution. The
95 code for `selfisher` was developed by modifying the R package `glmmTMB` (Brooks et al.,
96 2017) because `glmmTMB` already had the capabilities needed for fitting and analyzing binomial
97 mixed effect models. Previously, `glmmTMB` was developed by adapting the popular user
98 interface from `lme4` (Bates et al., 2015) and increasing the model flexibility and fitting
99 robustness by doing estimation with `TMB` (i.e. Template Model Builder, Kristensen et al., 2016).
100 Prior to the development of `glmmTMB`, `TMB` was developed based on the algorithm of AD Model
101 Builder (`ADMB`), which performs maximum likelihood estimation (MLE) in a fast and robust way
102 (Fournier et al., 2012; Miller, 2013). The algorithm is fast and robust because it has information
103 on the gradients of the likelihood surface via automatic differentiation. Additionally, `TMB`
104 improves robustness by providing binomial and beta-binomial likelihood functions that are
105 numerically stable even when probabilities are near zero or one. Thus, through inheritance,
106 `selfisher` has a flexible user interface with `lme4`-style syntax and robust `TMB` code
107 underlying the model estimation which is done using the same MLE algorithm as `ADMB`.

108 All models in `selfisher` involve comparing the catches from two compartments (e.g.
109 test vs control gear, gear 1 vs gear 2, or codend vs cover), which gives rise to data that can be
110 analysed as a binomial response, subject to the use of appropriate methods to allow for

111 overdispersion as described below. Due to technicalities of the underlying code, in `selfisher`
112 syntax, the binomial response must be specified as a proportion (i.e. proportion of individuals of
113 a given length in one compartment of one haul with respect to the total) and a total (i.e. total
114 number of individuals of a given length in either compartment in one haul), as we will
115 demonstrate in case studies; this is in contrast to other binomial regression methods that take a
116 two-column response variable. There are three main categories of experimental designs in
117 which each individual has two possible outcomes, i.e. studies that produce binomial data and
118 can be analyzed with `selfisher`. The categories are (1) a selective net inside an outer
119 nonselective small-mesh cover net (covered codend), (2) a selective net compared to a
120 nonselective net (paired gear), and (3) a comparison between two selective gears (catch
121 comparison). These can all be modeled using `selfisher` as we describe in sections below,
122 but first we describe some generalities.

123 Retention models

124 All three categories of analyses involve estimating a retention model; covered codend
125 and paired gears experiments allow one to estimate the absolute retention, i.e. retention out of
126 the population encountered by the gear, while in a catch comparison experiment the estimated
127 retention is relative to that of a baseline gear. Regardless, the mathematical formulation is
128 general. We use $r(l)$ to refer to the retention model throughout this text, but as we describe in
129 case studies below, it may depend on covariates other than length, l . See Table 1 for a
130 description of all notations. As in binomial generalized linear models (GLMs), retention models
131 use a link function to keep the retention probability in the range from zero to one. The most
132 common link is the "logit" (i.e. logistic), but other options include, "probit" (i.e. normal
133 probability ogive), "cloglog" (i.e. negative extreme value), "loglog" (i.e. extreme
134 value/Gompertz), or "Richards". The software default is the logit link. To fit retention model

135 shapes that are more diverse than the built-in link functions, it is possible to use a logit link with
136 more complex models such as polynomials (Holst & Reville, 2009) or smooth functions (Skalski &
137 Perez-Comas, 1993; Munro & Somerton, 2001; Fryer et al., 2003; Somerton et al., 2013) as
138 demonstrated in case studies C and D below.

139 Selectivity statistics I_{50} and SR

140 Two estimated summary statistics of interest are the length with 50% probability of
141 retention (I_{50}) and the selection range (SR, i.e. the width of the range of length classes with 25%
142 to 75% retention probabilities). Note that they only apply to models where retention probability
143 monotonically increases with length, such as in covered codend and paired gear studies. In
144 simple models with only length or size as a predictor of retention, then I_{50} and SR can be
145 extracted from a model using the function $L50SR()$. In more complex models, such as the
146 covered codend and paired gear case studies below, which involve additional covariates
147 besides length, there are multiple ways to extract I_{50} and SR estimates. The covered codend
148 case study demonstrates how to extract them algorithmically by finding the lengths that
149 correspond to retention probabilities 0.25, 0.50, and 0.75 for each given value of covariates in
150 the model. The paired gear example solves for I_{50} and SR mathematically using a model's
151 estimated coefficients. For either method, confidence intervals can be obtained by
152 bootstrapping.

153 Subsampled catches

154 Often, in cases of abundant catches, it may not be feasible to measure the length of
155 every individual that is caught and, in those cases, only a fraction of the catch may be
156 measured. This leads to additional statistical complexity in the analyses, but we have made

157 `selfisher` capable of handling subsampling in any model. Here we denote the approximate
158 fraction of individuals in compartment i , of haul h , and length class l that were sampled as $s_{i,h,l}$.
159 Although subsampling doesn't always depend on l , we have written the package to be flexible
160 enough to handle cases where each observed count (e.g. $n_{i,h,l}$) has a different subsampling
161 fraction. It is sufficient to include the ratio of subsampling fractions (i.e. the subsampling ratio) in
162 a model, rather than each compartment's fraction individually, $q_{h,l} = s_{i,h,l}/s_{j,h,l}$, assuming here
163 that i is the compartment being considered as "success" in the binomial context and j is the
164 alternative compartment. If raising factors were recorded in the data instead of subsampling
165 fractions, then care should be taken when calculating the subsampling ratio to account for the
166 fact that a raising factor is the inverse ($1/s_{i,h,l}$) of a subsampling fraction. Subsampling fractions
167 are between zero and one, while raising factors are greater than or equal to one.

168 The most general way to account for subsampling in a `selfisher` model is to specify
169 the subsampling ratio $q_{h,l}$ using the argument `qratio` in a call to the `selfisher` function. In
170 covered codend models with logit links (Millar, 1994) or catch comparison models with logit links
171 (Holst & Revill, 2009) an `offset` could be used equivalently, but using `qratio` will be more
172 broadly applicable because it can be used with any type of link including in paired gear models
173 in addition to the other types.

174 Three categories of experimental designs

175 Covered codend

176 One way of characterizing the selectivity of towed gear is to capture the individuals that
177 escape the net using a small-mesh cover, commonly known as the covered codend method.

178 Then, the statistical model has strong information on retention because the total number of
 179 individuals encountered in each length class is directly observed. In this category, we compare
 180 the number of individuals sampled in the cover in haul h with length l ($n_{1,h,l}$) and codend ($n_{2,h,l}$)
 181 by modeling the proportion $\frac{n_{1,h,l}}{n_{1,h,l} + n_{2,h,l}}$ as the probability of being retained and sampled in the
 182 codend divided by the probability of being retained and sampled in the codend or escaping to
 183 the cover and being sampled (Millar, 1994):

$$184 \quad \phi_{covered,h,l} = \frac{r(l)s_{1,h,l}}{r(l)s_{1,h,l} + (1-r(l))s_{2,h,l}} \text{ or more simply in terms of the subsampling ratio:}$$

$$185 \quad \phi_{covered,h,l} = \frac{r(l)q_{h,l}}{r(l)q_{h,l} + (1-r(l))}.$$

186 See Table 1 for definitions of all symbols.

187 Paired gear (where one gear is nonselective)

188 Another way to characterize the selectivity of a fishing gear is to acquire knowledge on
 189 the population available to be caught in each haul. Paired gear studies accomplish this by
 190 deploying a control gear or codend (besides the one whose selectivity is being measured) that
 191 has full retention for the species and lengths of interest, i.e. all individuals of the given species
 192 entering that gear are retained. We assume that there is a probability, p_h , of entering the test
 193 gear in haul h , given that the individual goes into either the test (t) or control (c) gear. The
 194 proportion of individuals observed in the test gear in haul h with length class l compared to the
 195 total number of individuals observed $\frac{n_{t,h,l}}{n_{t,h,l} + n_{c,h,l}}$ is modeled as the probability of entering the test
 196 net, being retained, and being sampled, divided by the probability of entering either net, being
 197 retained and being counted:

$$199 \quad \phi_{paired,h,l} = \frac{p_h s_{t,h,l} r(l)}{p_h s_{t,h,l} r(l) + (1-p_h) s_{c,h,l}}$$

200 It is convenient to write the model more simply in terms of the subsampling ratio:

$$201 \quad \Phi_{paired,h,l} = \frac{p_h q_{h,l} r(l)}{p_h q_{h,l} r(l) + 1 - p_h}$$

202 In paired gear studies, the ideal relative fishing power is 50% (i.e. $p=0.5$). If this is known
203 *a priori* then it is possible to fix p at 0.5 by specifying `pformula=~0` in the `selfisher` model.

204 If another value of p is known *a priori* due to differences in effort, such as swept areas, then that
205 can be specified using the `qratio` argument. For example, if the study has (haul-specific)
206 subsampling fractions `st` and `sc` as well as (haul-specific) swept areas of `at` and `ac` for the test
207 and control gear respectively, then one could use argument `qratio = at / ac * st / sc`
208 together with `pformula=~0` (e.g. Somerton et al., 2013).

209 Catch comparison (where both gears are selective)

210 Catch comparison studies compare the catches in two gears, both of which are
211 selective. Consequently, there is no direct information about the length distribution of the
212 population being fished and it is only possible to model the relative retention probability given
213 the population encountered during testing (Revill & Holst, 2004). In general, the relative
214 retention probability model can be of arbitrary shape and, for example, may not be monotone.

215 The response variable is the proportion of fish retained by one gear versus the other

$$216 \quad \frac{n_{1,h,l}}{n_{1,h,l} + n_{2,h,l}}$$

217 In a catch comparison study, the modeled proportion of the total catch retained and sampled in
218 gear 1 versus gear 2 is the probability of entering gear 1, being retained in gear 1, and being
219 sampled over the probability of entering, being retained, and being sampled in either gear. It is
220 always modeled with a logit link:

$$221 \quad \Phi_{compare,h,l} = \frac{p_h s_{1,h,l} r_1(l)}{p_h s_{1,h,l} r_1(l) + (1 - p_h) s_{2,h,l} r_2(l)}$$

222 where $r_1(l)$ and $r_2(l)$ are the absolute retention probabilities of the two gears, but it is not possible
223 to estimate them separately. Holst and Revill (2009) showed that the expected proportion can
224 be approximated with a polynomial with a logit link and an offset to account for any
225 subsampling. They showed that the relative fishing pressure p_h can be absorbed by the
226 intercept and that it may vary randomly between hauls and accounted for by including a random
227 intercept of haul as in a mixed effects model. In general, the relative retention model in a catch
228 comparison analysis can be formulated as either a polynomial (Holst & Revill, 2009) or a spline
229 (Miller, 2013); see the catch comparison case studies in Supplements C and D for a
230 demonstration of how this can be done with `selfisher`.

231 Handling extra variability

232 Overdispersion

233 Overdispersion is the presence of variability in the proportions that is in excess of the
234 variability specified under a binomial model as is typical for selectivity models. Overdispersion
235 can arise due to between-haul variability whereby the retention model varies from haul to haul.
236 To a lesser extent it can also arise due to within-haul variability due to schooling behavior that
237 violates the assumption that all fish behave independently. The accepted approaches for
238 including this variability are the use of bootstrapping (Millar, 1993) and the use of mixed effects
239 models (Fryer et al. 2003, Millar et al., 2004). Both of these methods are available in
240 `selfisher`. All of the case studies demonstrate how to bootstrap as described below. See the
241 covered codend and paired gear case studies (Supplements A and B) for examples of using
242 mixed effects in `selfisher`. In this text, we do not get into the details of random effects
243 because it is a large topic; however, note that in `selfisher`, they are implemented in the same

244 way as `glmmTMB` (Brooks et al., 2017) using the Laplace approximation, which is a standard
245 method commonly used in modern mixed modelling. Laplace approximation should work well
246 provided there is a sufficient sample size to inform the estimates for each haul (Ogden 2021),
247 but the sample sizes needed are an open area of research.

248 Quantify uncertainty by bootstrapping

249 For any statistical model, it is important to compare predictions from the model - and the
250 uncertainty around those predictions - with observed data to ensure that the model reasonably
251 represents the data. A bootstrapping procedure was developed by Millar (1993) to account for
252 variation between and within hauls and calculate appropriately wide confidence intervals. The
253 bootstrapping method can account for overdispersion in data due to variability among hauls (as
254 described above in the *Overdispersion* section); because of this, it is not necessary to include a
255 random effect of haul in models to be bootstrapped. The bootstrapping method is also
256 sometimes referred to as a “double bootstrap” in fishing gear selectivity literature, but this term
257 has another meaning in statistics (e.g. Kuk, 1989). This method first resamples the same
258 number of hauls from the observed set, with replacement. Then for each resampled haul, the
259 method resamples observed fish within the haul. Then it refits the model to the resampled data
260 set and the refit model is used to produce values such as predictions or parameter estimates. It
261 is typically repeated one thousand times or more. This bootstrapping method is implemented in
262 `selfisher` in a way that maintains all variables associated with each observed data point, not
263 just length class, e.g. sampling fractions or total catch. This is facilitated by specifying the `haul`
264 argument in the `selfisher` model fitting function; followed by a call to the `bootSel` function.

265 It is also possible to do resampling from pools of hauls, so that every bootstrapped
266 dataset has the same number of hauls in each pool as in the original data (Herrmann et al.,
267 2017). That is, if the original dataset has H_A hauls of type A and H_B hauls of type B, then it is

268 possible to bootstrap in a way such that simulated data sets have H_A hauls of type A and H_B
269 hauls of type B. This is done using the `pool` argument to the `selfisher` function. The inner
270 part of the bootstrapping method (resampling fish within hauls) is the same as in the regular
271 bootstrapping method. This is useful when hauls of the two gears or tactics being compared are
272 unpaired. See the gillnet case study in Supplement C for an example of specifying pools of
273 hauls.

274 Bootstrapping and mixed modeling

275 Mixed modelling is a formal method that takes into account possible sources of
276 variability in the data such as variation between hauls, enabling sound hypothesis testing and
277 model selection. However, fitting mixed models can be computationally intensive. Moreover, the
278 researcher is typically interested in obtaining overall selectivity predictions, rather than at the
279 haul level, because these are relevant to the selectivity applied to the fishery. These could be
280 obtained by taking a pseudo-Bayesian approach, that is, by simulating over all sources of
281 variability in the fitted mixed model, including random effects and parameter uncertainty.
282 However, this would require complex bespoke programming. A more straightforward alternative
283 would be to fit the best candidate model, but leaving out all random effects, and to perform a
284 bootstrap of this fixed-effects model by resampling the data in a way that emulates all additional
285 sources of variability (Millar, 2011). Bootstraps can then be used to obtain confidence intervals
286 for estimated quantities such as predicted retention curves or I_{50} and SR. See case studies A
287 and B for examples.

288 Case studies

289 A. Covered codend analyses of four codends catching haddock

290 This case study uses the haddock data from an experiment that employed the covered
291 codend method to investigate the selective performance of four codends (O'Neill et al., 2016).
292 The codends were made from netting materials with different twine bending stiffnesses and
293 mesh sizes.

294 We begin the case study by looking at just one gear type to demonstrate different link
295 functions that can be used and to show how to account for subsampling, which in this example
296 varies with length. The default link is the logistic, but we also consider the probit and Richards
297 curves, and a spline. The spline is fit using the `bs()` function from the `splines` package to
298 construct a basis for a polynomial spline with 3 degrees of freedom so that the model has an
299 intercept plus three nodes located based on quantiles of fish length. Penalized splines are not
300 yet implemented so the degrees of freedom are fixed. Having chosen a model, we bootstrap to
301 estimate 95% confidence interval for the proportion retained by the codend.

302 We then analyse all four gear types together and investigate the influence of length,
303 mesh size, bending stiffness and catch size. We assume the principle of geometric similarity (as
304 used by Tokai et al. (1996) to investigate grid selection). Geometric similarity is the assumption
305 that l_{50} and SR are proportional to mesh size and this corresponds to using a
306 $I(\text{length}/\text{meshsize})$ term in the `selfisher` formula interface for the retention model
307 (Baranov, 1948). We explore a number of models and choose the best fit using Bayesian
308 information criterion (BIC). When choosing the best model, we include a random effect of haul
309 to account for between-haul variation. As in the original publication, we show that selection is
310 dependent on all three parameters. Before bootstrapping, we drop the random effect of haul

311 from the best model because the bootstrapping method accounts for between-haul variability
312 and the random effect would slow it down considerably. We bootstrap to estimate confidence
313 intervals for the proportion retained by each gear and numerically solve for I_{50} and SR
314 dependent on covariates. See Supplement A for details and code.

315 Results were similar to the original manuscript, which also found a dependence of
316 haddock I_{50} on mesh size, twine bending stiffness, and total codend catch weight (Fig. 1) but
317 that SR is constant (O'Neill et al., 2016). We should not, however, expect the results to be
318 identical as there are many differences between the analyses. Originally, it was assumed that
319 the slope and intercept of the logistic link functions vary randomly from haul to haul and that I_{50}
320 and $\log(\text{SR})$ mean selection curve were linearly dependent on the explanatory variables using
321 separate models. Here, we have used a single model which assumed geometric similarity for
322 mesh size and that overall retention is related to the explanatory variables.

323 B. Paired gear analyses of codend selectivity dependent on 324 mesh size

325 This case study draws on a subset of data from the German research project CRANNET
326 (Santos et al., 2018). The experimental method consisted of fishing with two identical beam
327 trawls, simultaneously and in parallel on the same shrimp population. One of the trawls had a
328 small-mesh (11 mm) control codend with very limited selectivity, assumed to be nonselective on
329 the range of shrimp lengths available to the trawl. The second trawl had a test codend. The
330 subset of data analyzed here consists of catch data from 87 hauls, during which 13 diamond-
331 mesh codends varying in mesh size ranging from 19.1 mm to 36.3 mm were tested. The goal
332 was to model I_{50} and SR as a function of mesh size, and to quantify any effect of two additional
333 haul covariates, sea state and catch weight.

334 The statistical modeling of selectivity begins with a mixed model to formally assess the
335 effect of mesh size, sea state and catch weight, while controlling for additional random variation
336 among hauls using a random intercept of haul. *A priori*, we assumed geometric similarity for the
337 mesh size (ms) covariate as in case study A above, by including the term $I(\text{length}/\text{ms})$ in
338 models. Sea state and test catch weight each had one value per haul in the data set. So that,
339 for a given paired haul of the test and control nets, the same value of sea state was used as a
340 predictor of retention probability of the test net for all length classes. The default model was
341 compared to several others and found to be preferred (using BIC), and neither sea state nor
342 catch weight had a significant effect.

343 Having chosen geometric similarity (with respect to mesh size) as the preferred model,
344 this model was refitted without random effects so as to estimate size selectivity at the population
345 level. Bootstrapping was used to obtain appropriate confidence intervals on I_{50} and SR for any
346 given mesh size. In addition, this case study demonstrates the use of `psplit=TRUE` (unequal
347 fishing power of the paired codends), the use of sampling ratios, use of the `inits()` function to
348 specify good starting values (because without it some models converged to local minima that
349 did not make any sense), and shows how I_{50} and SR can be obtained directly from the model
350 fitted by `selfisher`. See Supplement B for details and code.

351 Estimates of I_{50} and SR dependent on mesh size were very close to those from the
352 original analysis (Fig. 2). The `selfisher` model predicted higher values of SR, with larger
353 uncertainty than Santos et al. (2018). This is a plausible result considering the different model
354 structures applied, and the large variation of the by-haul estimates.

355 C. Catch comparison analyses of unpaired hauls of gillnets 356 avoiding an unwanted crab

357 This example deals with data from an experiment originally published by Savina et al.
358 (2017). Two soak tactics (12h at day and 12h at night) were compared in the Danish gillnet
359 plaice fishery to estimate the role that the choice of a soak tactic plays in the catch efficiency of
360 both target and unwanted species. This is a subset of the original dataset (one species, two
361 soak tactics) where we are looking at the unwanted invertebrate edible crab (*Cancer pagurus*).

362 We use the method developed by Herrmann et al. (2017) which was developed for
363 assessing the effect of changing the gear design on the relative length-dependent catch
364 efficiency. This example is representative of experimental fishing where the catch data obtained
365 for two gears or tactics were not collected in pairs, and can allow for a different number of
366 deployments.

367 This case study is a typical model for catch comparison of multiple haul data without
368 subsampling using a spline. To get confidence intervals on predictions, we bootstrapped from
369 two pools according to tactic (night vs day) using the argument `pool=tactic` to the `bootSel`
370 function. See Supplement C for details and code.

371 The estimated catch comparison curves and relative confidence intervals were very
372 similar to the original analysis (Fig. 3). The catch comparison curves properly reflected the trend
373 in the experimental points. The experimental rates were subject to increasing binomial noise
374 outside the length classes representing the main bulk of the catches. The results for edible crab
375 showed significantly higher catches at night than at day with no strong indication of a length
376 dependency.

377

378 D. Catch comparison analyses of paired hauls of *Nephrops* twin- 379 rigged trawls

380 The example is based on the data from Melli et al. (2018). An anterior gear modification,
381 namely the counter-herding device FLEXSELECT, was tested in a twin-rig configuration, where
382 two identical trawls were towed in parallel. One trawl was equipped with FLEXSELECT, referred
383 to as the test trawl, while the other worked as baseline. The aim of the study was to determine if
384 FLEXSELECT could reduce the fish bycatch in a *Nephrops*-directed fishery. The data used in the
385 example are from haddock, which was found to be strongly affected by the counter-herding
386 device.

387 Following the steps of the published paper, we conducted a catch comparison analysis,
388 modeling the relative retention as a 4th-order polynomial. In addition, we used a spline with 4
389 degrees of freedom using the `splines` package and performed model selection to determine if
390 it fitted the data better. Considering that part of the hauls were conducted in day-time and part in
391 night-time, “time” was included in the model as an explanatory variable to determine if the
392 length-based efficiency of FLEXSELECT presented diel differences. We predicted both catch
393 comparison rates and catch ratio with bootstrapped confidence intervals using the `predict`
394 and `bootSel` functions from `selfisher`. See Supplement D for details and code.

395 Model selection using AIC indicated that a spline with five degrees of freedom was only
396 slightly more parsimonious than a fourth-order polynomial (0.4 deltaAIC). Because the
397 polynomial has fewer parameters with nearly equal AIC (df=12 compared to df=10) we plot
398 results from the polynomial here to show that results from `selfisher` were nearly equal to the
399 analyses from the original publication which used `SELNET` (Fig. 4.).

400

401

402 Discussion

403 We have introduced an open source R package for estimating fishing gear selectivity of
404 both towed and passive gear, making it easier for anyone to analyze fishing gear selectivity data
405 without writing extensive amounts of code. We have demonstrated its broad applicability in four
406 case studies spanning a range of experimental designs. The case studies have shown that
407 results from `selfisher` are comparable to previously published results and that `selfisher`
408 is more flexible than some methods (e.g. a single model to quantify the effect of changing mesh
409 size on I_{50} and SR). Some of the features of `selfisher` that were demonstrated in the case
410 studies are summarized in Table 2. The case studies aim to demonstrate best practices based
411 on current knowledge. However, this is an active area of research and with a new powerful
412 model fitting tool, best practices may change. Even with (or especially with) a powerful tool,
413 analyses require careful thought and checking of results. For example, in complicated models
414 such as paired gear models which contain two submodels (retention and relative fishing power),
415 it may be necessary to be cautious about identifiability of parameters and local optima
416 encountered during maximum likelihood estimation, but better starting values help avoid those
417 issues as demonstrated in Supplement B (Bolker et al., 2013). If some parts of a model are not
418 identifiable, then the summary output will contain “NA”s. The `TMBhelper` package (Thorson,
419 2020) contains a function to identify which parts of a `selfisher` model (e.g. `mod`) may be
420 problematic by typing `check_estimability(mod$obj)`.

421 We have several advancements for the package either planned for the future or already
422 underway. Thus, we recommend installing the latest version and keeping up to date with
423 advancements by following the guidance in Appendix 1. We plan to add functions to calculate
424 discard ratio indices and indicator functions (e.g. Wienbeck et al., 2014; Santos et al., 2016;
425 Veiga-Malta et al., 2019). We may add a function to facilitate model averaging, although it is

426 already possible to piece this functionality together with the existing features. We have not tried
427 to fit structured non-monotonic curves (e.g. bell-shaped curves of gillnet absolute selectivity
428 based on geometric similarity, Baranov 1948) with `selfisher`, but we will explore this
429 possibility in the future. We will investigate how to choose starting values of parameters in
430 models that have Richards link, to increase robustness. We plan to implement a general method
431 to extract I_{50} and SR from complex models as demonstrated in case studies A and B. To handle
432 overdispersion more elegantly, we plan to add the option of having a beta-binomial distribution
433 for the response (Miller, 2013). We are already in the process of developing a Shiny app, which
434 will facilitate simple standard analyses without the need for writing code; this will help to bridge
435 the gap for scientists or managers with extensive experience in gear development but little
436 experience with R. As an open source package, code developers are encouraged to contribute
437 improvements through GitHub such as those listed here.

438 Having access to a free and open source software should benefit this field of research in
439 several ways. It allows researchers to share code and thereby foster a community for discussion
440 and repeatability. The free nature of the software will enable researchers and managers with
441 limited budgets - such as those in developing countries - to perform analyses themselves. It
442 gives statistical methods of retention modelling a way into classrooms containing the next
443 generation of fisheries scientists who are already learning modern regression methods as part
444 of a general scientific curriculum.

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452 Contributors' statement

453 Conceptualization: MB, LK, JF; Methodology, MB, RM; Software, MB, TVM; Validation: MB,
454 VM, TVM, ES, JS, BON; Data Curation: MB, VM, ES, JS, BON; Writing - Original Draft: MB, VM,
455 TVM, ES, RM, JS, BON; Writing - Review & Editing: MB, VM, TVM, ES, RM, JS, BON, JF
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463 Flexible solutions for the Nephrops Fisheries (33113-I-16-068).

464 Data Availability Statement

465 Data used in the case studies is available as part of the `selfisher` package on GitHub. See
466 the *Installation* instructions in Appendix 1 as well as the online supplementary material.

467

References

- 468 Baranov, F. I. (1948). Theory and assessment of fishing gear. In Theory of fishing with gillnets.
469 Chap. 7. Pishchepromizdat, Moscow. (Translation from Russian by Ontario Dept of
470 Lands, Maple, Ont., 45 pp.)
- 471 Bates, D., Mächler, M., Bolker, B., & Walker, S. (2015). Fitting linear mixed-effects models using
472 lme4. *Journal of Statistical Software*, 67(1), 1–48. <http://dx.doi.org/10.18637/jss.v067.i01>
- 473 Bolker, B. M., Gardner, B., Maunder, M., Berg, C.W., Brooks, M., Comita, L., Crone, E.,
474 Cubaynes, S., Davies, T., de Valpine, P., Ford, J., Gimenez, O., Kéry, M., Kim, E.J.,
475 Lennert-Cody, C., Magnusson, A., Martell, S., Nash, J., Nielsen, A., Regetz, J., Skaug,
476 H. and Zipkin, E. (2013), Strategies for fitting nonlinear ecological models in R, AD
477 Model Builder, and BUGS. *Methods Ecol Evol*, 4: 501-512. [https://doi.org/10.1111/2041-](https://doi.org/10.1111/2041-210X.12044)
478 [210X.12044](https://doi.org/10.1111/2041-210X.12044)
- 479 Brooks, M. E., Kristensen, K., van Benthem, K. J., Magnusson, A., Berg, C. W., Nielsen, A.,
480 Skaug, H.J., Machler, M., & Bolker, B. M. (2017). glmmTMB balances speed and
481 flexibility among packages for zero-inflated generalized linear mixed modeling. *The R*
482 *journal*, 9(2), 378-400. <https://doi.org/10.3929/ethz-b-000240890>
- 483 Erickson, D. L., Perez-Comas, J. A., Pikitch, E. K., & Wallace, J. R. (1996). Effects of catch size
484 and codend type on the escapement of walleye pollock (*Theragra chalcogramma*) from
485 pelagic trawls. *Fisheries Research*, 28(2), 179-196. [https://doi.org/10.1016/0165-](https://doi.org/10.1016/0165-7836(96)00497-3)
486 [7836\(96\)00497-3](https://doi.org/10.1016/0165-7836(96)00497-3)
- 487 Ferro, R. S. T., Özbilgin, H., & Breen, M. (2008). The potential for optimizing yield from a
488 haddock trawl fishery using seasonal changes in selectivity, population structure and fish

- 489 condition. *Fisheries research*, 94(2), 151-159.
490 <https://doi.org/10.1016/j.fishres.2008.08.018>
- 491 Fournier, D. A., Skaug, H. J., Ancheta, J., Ianelli, J., Magnusson, A., Maunder, M. N., Nielsen,
492 A., & Sibert, J. (2012). AD Model Builder: using automatic differentiation for statistical
493 inference of highly parameterized complex nonlinear models. *Optimization Methods and*
494 *Software*, 27(2), 233-249. <https://doi.org/10.1080/10556788.2011.597854>
- 495 Fryer, R. J. (1991). A model of between-haul variation in selectivity. *ICES Journal of Marine*
496 *Science*, 48(3), 281-290. <https://doi.org/10.1093/icesjms/48.3.281>
- 497 Fryer, R. J., Zuur, A. F., & Graham, N. (2003). Using mixed models to combine smooth size-
498 selection and catch-comparison curves over hauls. *Canadian Journal of Fisheries and*
499 *Aquatic Sciences*, 60(4), 448-459. <https://doi.org/10.1139/f03-029>
- 500 He, P. (1993). Swimming speeds of marine fish in relation to fishing gears. *ICES Mar. Sci.*
501 *Symp*, 196, 183-189.
- 502 Herrmann, B., Sistiaga, M., Rindahl, L., & Tatone, I. (2017). Estimation of the effect of gear
503 design changes on catch efficiency: methodology and a case study for a Spanish
504 longline fishery targeting hake (*Merluccius merluccius*). *Fisheries Research*, 185, 153-
505 160. <https://doi.org/10.1016/j.fishres.2016.09.013>
- 506 Holst, R., & Revill, A. (2009). A simple statistical method for catch comparison studies. *Fisheries*
507 *Research*, 95(2-3), 254-259. <https://doi.org/10.1016/j.fishres.2008.09.027>
- 508 [Kenelly, S.J., & Broadhurst, M.K. 2021. A review of bycatch reduction in demersal fish trawls.](#)
509 [Reviews in Fish Biology and Fisheries, https://doi.org/10.1007/s11160-021-09644-0](#)

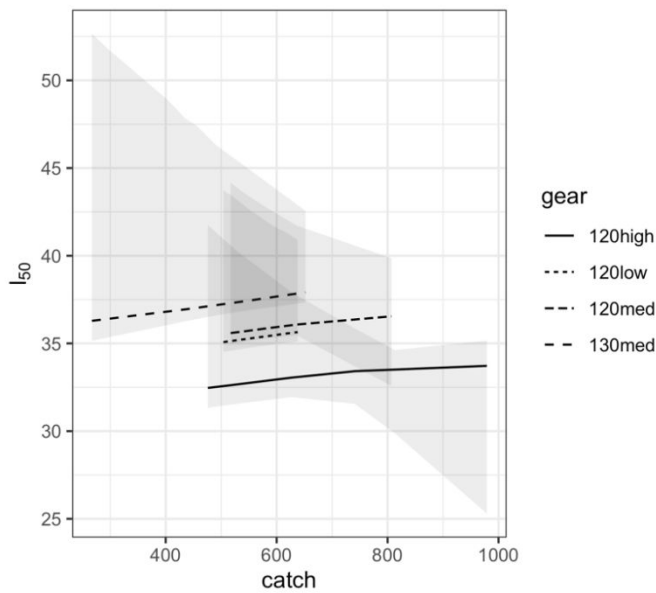
- 510 Kristensen, K., Nielsen, A., Berg, C.W., Skaug, H., Bell, B.M. (2016). TMB: Automatic
511 Differentiation and Laplace Approximation. *Journal of Statistical Software*, 70(5), 1-21.
512 <http://dx.doi.org/10.18637/jss.v070.i05>
- 513 Kuk, A. Y. (1989). Double bootstrap estimation of variance under systematic sampling with
514 probability proportional to size. *Journal of Statistical Computation and Simulation*, 31(2),
515 73-82.
- 516 Larsen, R. B., Herrmann, B., Sistiaga, M., Brinkhof, J., & Grimaldo, E. (2018). Bycatch reduction
517 in the Norwegian deep-water shrimp (*Pandalus borealis*) fishery with a double grid
518 selection system. *Fisheries Research*, 208, 267-273.
519 <https://doi.org/10.1016/j.fishres.2018.08.007>
- 520 Melli, V., Karlsen, J. D., Feekings, J. P., Herrmann, B., & Krag, L. A. (2018). FLEXSELECT:
521 counter-herding device to reduce bycatch in crustacean trawl fisheries. *Canadian*
522 *Journal of Fisheries and Aquatic Sciences*, 75(6), 850-860. [https://doi.org/10.1139/cjfas-](https://doi.org/10.1139/cjfas-2017-0226)
523 [2017-0226](https://doi.org/10.1139/cjfas-2017-0226)
- 524 Melli, V., Krag, L. A., Herrmann, B., & Karlsen, J. D. (2019). Can active behaviour stimulators
525 improve fish separation from Nephrops (*Nephrops norvegicus*) in a horizontally divided
526 trawl codend?. *Fisheries Research*, 211, 282-290.
527 <https://doi.org/10.1016/j.fishres.2018.11.027>
- 528 Millar, R. B. (1993). Incorporation of between-haul variation using bootstrapping and
529 nonparametric estimation of selection curves. *Fishery Bulletin*, 91, 564-572.
- 530 Millar, R. B. (1994). Sampling from trawl gears used in sized selectivity experiments. *ICES*
531 *Journal of Marine Science*, 51(3), 293-298. <https://doi.org/10.1006/jmsc.1994.1030>

- 532 Millar, R. B. (2011). Maximum likelihood estimation and inference: with examples in R, SAS and
533 ADMB. Wiley, London. 357 p. ISBN: 978-0-470-09482-2
- 534 Millar, R. B., & Fryer, R. J. (1999). Estimating the size-selection curves of towed gears, traps,
535 nets and hooks. *Reviews in Fish Biology and Fisheries*, 9(1), 89-116.
536 <https://doi.org/10.1023/A:1008838220001>
- 537 Millar, R. B., Broadhurst, M. K., & Macbeth, W. G. (2004). Modelling between-haul variability in
538 the size selectivity of trawls. *Fisheries Research*, 67(2), 171-181.
539 <https://doi.org/10.1016/j.fishres.2003.09.040>
- 540 Miller, T. (2013). A comparison of hierarchical models for relative catch efficiency based on
541 paired-gear data for US Northwest Atlantic fish stocks. *Can. J. Fish. Aquat. Sci.* 70(9),
542 1306–1316. <https://doi.org/10.1139/cjfas-2013-0136>
- 543 Munro, P. T., & Somerton, D. A. (2001). Maximum likelihood and non-parametric methods for
544 estimating trawl footrope selectivity. *ICES Journal of Marine Science*, 58(1), 220-229.
545 <https://doi.org/10.1006/jmsc.2000.1004>
- 546 O'Neill F.G., Kynoch R.J., Blackadder L., Fryer, R.J., Eryaşar A.R., Notti, E., & Sala A. (2016).
547 The influence of twine tenacity, thickness and bending stiffness on codend selectivity.
548 *Fisheries Research*, 176, 94–99. <https://doi.org/10.1016/j.fishres.2015.12.012>
- 549 O'Neill, F. G., & Kynoch, R. J. (1996). The effect of cover mesh size and cod-end catch size on
550 cod-end selectivity. *Fisheries Research*, 28(3), 291-303. [https://doi.org/10.1016/0165-](https://doi.org/10.1016/0165-7836(96)00501-2)
551 [7836\(96\)00501-2](https://doi.org/10.1016/0165-7836(96)00501-2)
- 552 Ogden, H. (2021). On the error in Laplace approximations of high-dimensional integrals. *Stat*,
553 10(1), e380. <https://doi.org/10.1002/sta4.380>

- 554 Özbilgin, H., Tosunoğlu, Z., Tokaç, A., & Metin, G. (2007). Seasonal variation in the trawl
555 codend selectivity of picarel (*Spicara smaris*). *ICES Journal of Marine Science*, 64(8),
556 1569-1572. <https://doi.org/10.1093/icesjms/fsm115>
- 557 Pérez Roda, M.A., Gilman, E., Huntington, T., Kennelly, S.J., Suuronen, P., Chaloupka, M.,
558 Medley, P. (2019) A third assessment of global marine fisheries discards. FAO Fisheries
559 and Aquaculture Technical Paper No. 633, FAO, Rome, 78pp
- 560 R Core Team (2020). R: A language and environment for statistical computing. R Foundation for
561 Statistical Computing, Vienna, Austria. URL <http://www.R-project.org/>.
- 562 Revill, A.S., & Holst, R. (2004) Reducing discards of North Sea brown shrimp (*C. crangon*) by
563 trawl modification. *Fisheries Research*, 68, 113-122.
564 <https://doi.org/10.1016/j.fishres.2004.02.001>
- 565 Ryer, C. H., & Barnett, L. A. (2006). Influence of illumination and temperature upon flatfish
566 reactivity and herding behavior: potential implications for trawl capture efficiency.
567 *Fisheries Research*, 81(2-3), 242-250. <https://doi.org/10.1016/j.fishres.2006.07.001>
- 568 Santos, J., Herrmann, B., Mieske, B., Stepputtis, D., Krumme, U., & Nilsson, H. (2016).
569 Reducing flatfish bycatch in roundfish fisheries. *Fisheries Research*, 184, 64-73.
570 <https://doi.org/10.1016/j.fishres.2015.08.025>
- 571 Santos, J., Herrmann, B., Stepputtis, D., Günther, C., Limmer, B., Mieske, B., and Schultz, S. et
572 al. (2018). Predictive framework for codend size selection of brown shrimp (*Crangon*
573 *crangon*) in the North Sea beam-trawl fishery. *PLoS ONE* 13(7): e0200464.
574 <https://doi.org/10.1371/journal.pone.0200464>

- 575 Savina, E., Krag, L.A., Frandsen, R.P., & Madsen, N. (2017). Effect of fisher's soak tactic on
576 catch pattern in the Danish gillnet plaice fishery. *Fisheries Research*, 196, 56-65.
577 <https://doi.org/10.1016/j.fishres.2017.08.009>
- 578 Skalski, J. R., & Perez-Comas, J. (1993). Nonparametric maximum likelihood estimation of
579 mesh size selectivity. *Fisheries research*, 18(3-4), 321-334. [https://doi.org/10.1016/0165-](https://doi.org/10.1016/0165-7836(93)90160-9)
580 [7836\(93\)90160-9](https://doi.org/10.1016/0165-7836(93)90160-9)
- 581 Somerton, David A., Weinberg, Kenneth L., Goodman, Scott E., & Chen, Yong. (2013)
582 Catchability of snow crab (*Chionoecetes opilio*) by the eastern Bering Sea bottom trawl
583 survey estimated using a catch comparison experiment. *Canadian Journal of Fisheries*
584 *and Aquatic Sciences*, 70(12), 1699-1708. <https://doi.org/10.1139/cjfas-2013-0100>
- 585 Suuronen, P., & Millar, R. B. (1992). Size selectivity of diamond and square mesh codends in
586 pelagic herring trawls: only small herring will notice the difference. *Canadian Journal of*
587 *Fisheries and Aquatic Sciences*, 49(10), 2104-2117. <https://doi.org/10.1139/f92-234>
- 588 Thorson, J. (2020). TMBhelper: Package for basic helper functions that are not worth putting in
589 a specialized contributed package. R package version 1.3.0.
- 590 Tokai, T., Omoto, S., Sato, R., & Matuda, K. (1996). A method of determining selectivity curve of
591 separator grid. *Fisheries Research*, 27(1-3), 51-60. [https://doi.org/10.1016/0165-](https://doi.org/10.1016/0165-7836(95)00471-8)
592 [7836\(95\)00471-8](https://doi.org/10.1016/0165-7836(95)00471-8)
- 593 Veiga-Malta, T., Feekings, J. P., Frandsen, R. P., Herrmann, B., & Krag, L. A. (2020). Testing a
594 size sorting grid in the brown shrimp (*Crangon Crangon* Linnaeus, 1758) beam trawl
595 fishery. *Fisheries Research*, 231, 105716. <https://doi.org/10.1016/j.fishres.2020.105716>

- 596 Veiga-Malta, T., Feekings, J., Herrmann, B., & Krag, L. A. (2019). Industry-led fishing gear
597 development: Can it facilitate the process?. *Ocean & Coastal Management*, 177, 148-
598 155. <https://doi.org/10.1016/j.ocecoaman.2019.05.009>
- 599 Walsh, S. J., & Hickey, W. M. (1993). Behavioural reactions of demersal fish to bottom trawls at
600 various light conditions. *ICES Mar. Sci. Symp*, 196, 68-76.
- 601 Wienbeck, H., Herrmann, B., Feekings, J. P., Stepputtis, D., & Moderhak, W. (2014). A
602 comparative analysis of legislated and modified Baltic Sea trawl codends for
603 simultaneously improving the size selection of cod (*Gadus morhua*) and plaice
604 (*Pleuronectes platessa*). *Fisheries Research*, 150, 28-37.
605 <https://doi.org/10.1016/j.fishres.2013.10.007>
- 606 Wileman, D., Ferro, R. S. T., Fonteyne, R., & Millar, R. B. (1996). Manual of methods of
607 measuring the selectivity of towed fishing gears. *ICES Cooperative Research Report*,
608 215.
609

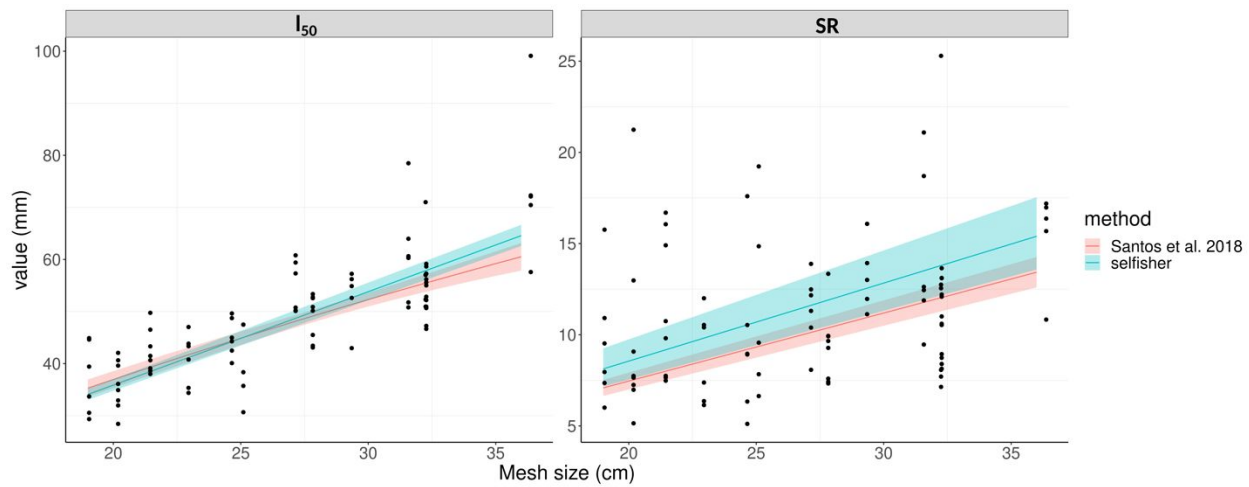
610 **Figures**

611

612 Figure 1. Estimated l_{50} as a continuous function of catch weight (kg) for four types of gear using
613 a single `selfisher` model. The lines represent the median prediction from 1000 bootstraps of
614 the `selfisher` model, while the bands show the 2.5 and 97.5 percentiles.

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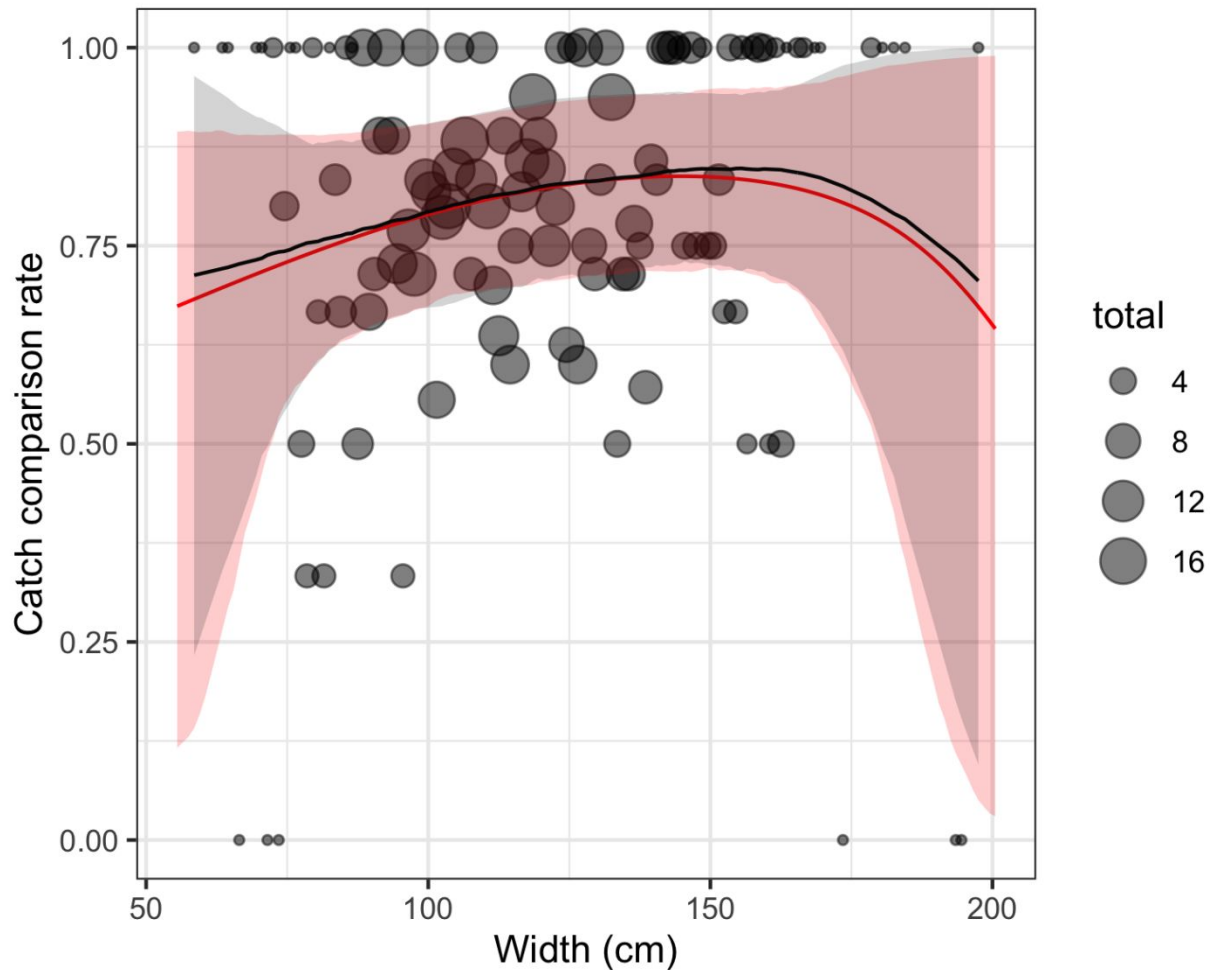
616



617

618 Figure 2. Estimated I_{50} and SR dependent on mesh size. The black points in the figure represent
 619 values of I_{50} and SR estimated at the haul level in the original study (Santos et al. 2018) and the
 620 red line and ribbon represent the analysis which combined those haul-specific estimates. The
 621 green line represents the median prediction from 1000 bootstraps of the `selfisher` model,
 622 while the band shows the 2.5 and 97.5 percentiles.

623



624

625 Figure 3. Catch comparison rate of night vs day fishing tactics for an unwanted crab in a gillnet

626 fishery. The black line and grey ribbon represent the median, 2.5, and 97.5 percentiles from

627 1000 bootstraps from a `selfisher` model using a spline with three degrees of freedom. The628 red line and ribbon are predictions from `SELNET` using model averaging. Circles are total

629 observations from the raw data, the size of which represents the total number of crabs caught

630 with either tactic.

631

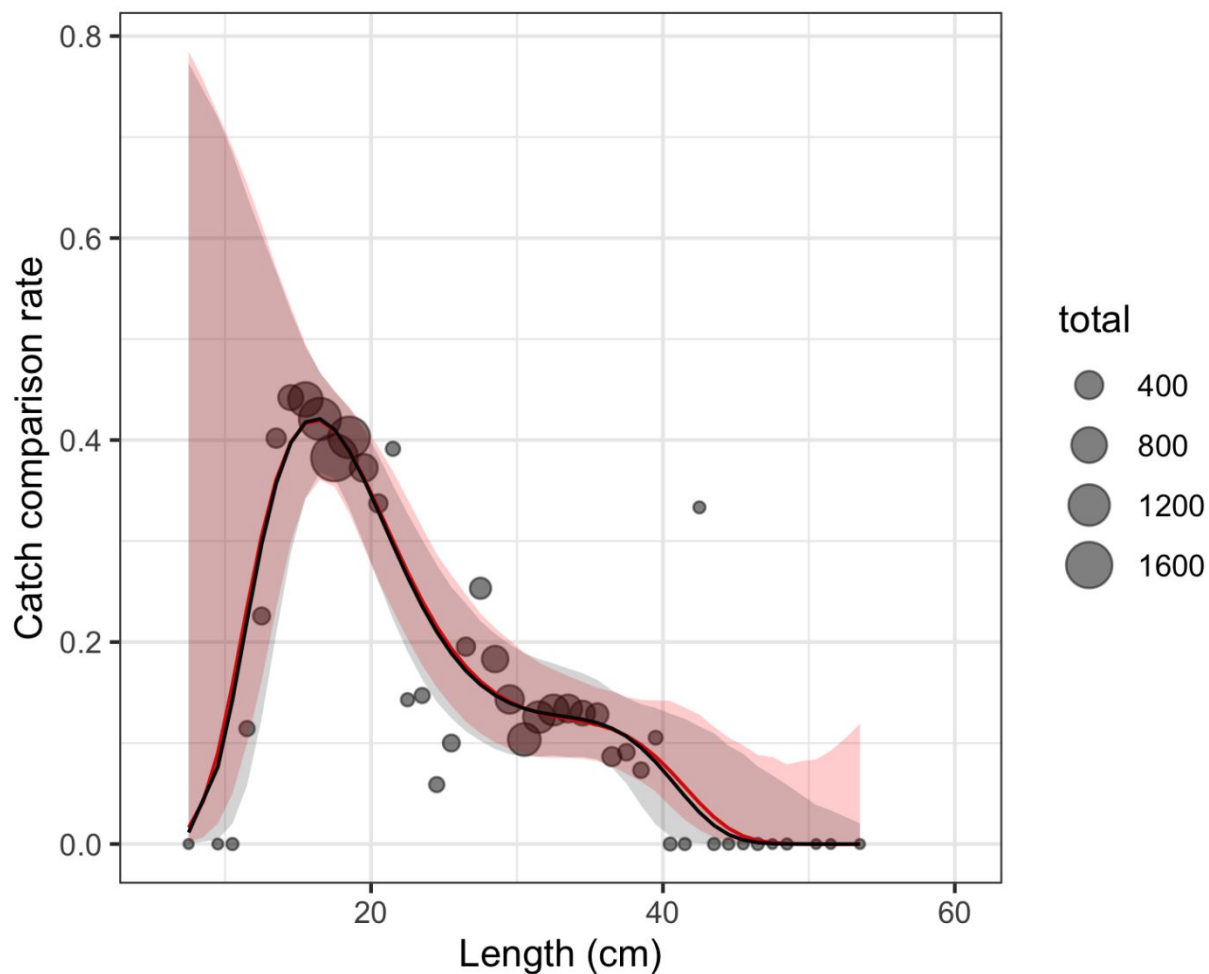


Figure 4. Catch comparison rates. Lines and ribbons are estimates and confidence intervals from `SELNET` in red using model averaging and from `selfisher` in black using a 5th order polynomial. Circles are observations from the raw data, aggregated over hauls, the size of which represents the number of haddock counted.

638 Table 1. Symbols used in the text.

Symbol	Meaning
$n_{i,h,l}$	Number of fish sampled in compartment i , haul h , in length class l
$s_{i,h,l}$	Proportion of fish sampled in compartment i , haul h , in length class l
$q_{h,l} = s_{i,h,l}/s_{j,h,l}$	Subsampling ratio for compartment i vs j in haul h , in length class l
$r(l)$	Retention probability as a function of length l
p_h	Relative fishing power of test vs control gear in haul h
$\phi_{paired,h,l}$	Probability model in a paired gear model
$\phi_{compare,h,l}$	Probability model in a catch comparison model
$\phi_{covered,h,l}$	Probability model in a covered codend model

639

640 Table 2. A list of `selfisher` features demonstrated in the four case studies with code
 641 in Supplements A through D.

Feature	Demonstrated in case studies
Random effects	A, B
Fixed effects other than length or size	A, B
Splines	C, D
Residual plot	A
Specifying initial values	B
Model selection via information criteria	A, B, C, D
Link functions other than default logit	A

642